

**Loss Models Prelims for Actuarial Students**  
**Monday, January 13, 2025, 9:00 AM - 1:00 PM**

**Instructions:**

1. There are five (5) equally-weighted questions and you are to answer all five.
2. Hand-held calculators are permitted.
3. Please provide details of your workings in the appropriate spaces provided; partial points will be granted.
4. Please write legibly. Points will be deducted for incoherent, incorrect, and/or irrelevant statements.

**Question No. 1:**

Let  $X$  and  $Y$  denote two (non-negative) random losses, and assume that their survival functions are given, respectively, by

$$S_X(x) = \begin{cases} 0.3, & \text{for all } 0 \leq x < 2, \\ 0.1, & \text{for all } 2 \leq x < 4, \\ 0, & \text{for all } x \geq 4, \end{cases}$$

and

$$S_Y(y) = \left( \frac{2}{2+y} \right)^3, \quad \text{for all } y \geq 0.$$

- (a) Calculate the mean and variance of  $X$ .
- (b) Calculate the mean and variance of  $Y$ .
- (c) Calculate VaR and CVaR at 80% level for both  $X$  and  $Y$ .

**Question No. 2:**

The probability generating function (pgf) of a random variable  $X$  is defined to be

$$P_X(z) = \mathbb{E}[z^X].$$

- (a) Prove, by recursion, that the pgf generates the probabilities for a discrete random variable  $N$ , i.e., prove that

$$p_n = \Pr(N = n) = \left( \frac{1}{n!} \right) P_N^{(n)}(0) = \left( \frac{1}{n!} \right) \frac{d^n}{dz^n} P_N(z) \Big|_{z=0},$$

for  $n = 0, 1, 2, \dots$

- (b) Use the result in (a) to describe the distribution of  $N$  with pgf

$$P_N(z) = \frac{z}{5}(2 + 3z^2).$$

- (c) Consider the aggregate loss models defined by

$$S_N = X_1 + X_2 + \cdots + X_N,$$

where  $X_1, X_2, \dots$  are identically, independent random variables describing claim severities, and  $N$  is a discrete random variable describing claim frequencies. All random variables are independent. Prove that the pgf of  $S_N$  can be expressed as

$$P_{S_N}(z) = P_N(P_X(z)).$$

### Question No. 3:

Consider a collective risk model

$$S = \sum_{i=1}^N X_i,$$

in which  $N$  follows a Poisson distribution with mean 10. Regarding the loss  $X$ , it is known that

$$\mathbb{E}[X] = 40,$$

$$\mathbb{E}[X \wedge 20] = 15,$$

$$\Pr(X > 20) = 0.8, \text{ and}$$

$$\mathbb{E}[X^2 | X > 20] = 4,000.$$

Now, assume that a deductible of  $d = 20$  (per loss) is applied to all policies, resulting in a new collective risk model for the aggregate payment

$$\tilde{S} = \sum_{i=1}^{\tilde{N}} \tilde{X}_i,$$

in which  $\tilde{X}_i := (X_i - d) | X_i > d$ , and  $\tilde{N}$  counts the number of *payment* events.

- (a) Show that  $\tilde{N}$  follows a Poisson distribution and compute its mean.  
 (b) Calculate the mean and variance of  $\tilde{S}$ .

### Question No. 4:

Individual ground-up loss amount  $X$  follows a two-parameter Weibull distribution with  $\tau = 1/3$  and mean 60. See Appendix for parameterization of Weibull in terms of  $\tau$  and  $\theta$ . An insurance policy on  $X$  has a deductible amount of 5 and a policy limit of 120 per loss.

Assume loss amount increased due to inflation by 5% uniformly.

- (a) Determine the value of the parameter  $\theta$ .

- (b) Calculate the expected value of claims per loss before the inflation. You may leave your answer in terms of the incomplete gamma function.
- (c) Calculate the expected value of claims per loss after the inflation. You may leave your answer in terms of the incomplete gamma function.
- (d) Calculate the variance of claims per loss after the inflation. You may leave your answer in terms of the incomplete gamma function.

**Question No. 5:**

Claim frequency  $N$  is being modeled as a zero-inflated Poisson expressed as follows:

$$\begin{aligned} p_0^m &= \pi + (1 - \pi)p_0 \\ p_k^m &= (1 - \pi)p_k, \quad \text{for } k = 1, 2, \dots \end{aligned}$$

where  $p_0 = e^{-\lambda}$  and  $p_k = e^{-\lambda} \frac{\lambda^k}{k!}$  for  $k = 1, 2, \dots$  are the probabilities of an ordinary Poisson random variable.

Now consider the random variable  $X$  defined by

$$X = Z \cdot Y$$

where  $Z \sim \text{Bernoulli}(1 - \pi)$ ,  $Y \sim \mathcal{PN}(\lambda)$ , and  $Z, Y$  are independent random variables.

- (a) Prove that  $N$  and  $X$  are equivalent random variables by showing that the probability mass function of  $X$  is exactly the same as that of  $N$ , as given above.
- (b) Use this result to show that the mean of  $N$  is

$$\mathbb{E}[N] = \mu = (1 - \pi)\lambda$$

and its variance is

$$\text{Var}(N) = \mu + \frac{\pi}{1 - \pi} \mu^2.$$

- (c) A variation of the method of moments is to match  $\mathbb{E}[N]$  with sample mean and  $\text{Var}(N)$  with sample variance. Given observed data  $x_1, x_2, \dots, x_n$ , use this suggested method of moments to show that resulting moment estimators of  $\pi$  and  $\lambda$  respectively can be expressed as

$$\tilde{\pi}_{\text{ME}} = \frac{s^2 - \bar{x}}{\bar{x}^2 + (s^2 - \bar{x})}$$

and

$$\tilde{\lambda}_{\text{ME}} = \bar{x} + \left( \frac{s^2}{\bar{x}} - 1 \right),$$

where

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i \quad \text{and} \quad s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$

are the respective sample mean and sample variance.

- (d) Explain, in words, why zero-inflated Poisson is more realistic and applicable to insurance losses than ordinary Poisson.

— end —

## APPENDIX

A random variable  $X$  is said to have a two-parameter Weibull distribution if its density has the form

$$f(x) = \frac{1}{x} \tau (x/\theta)^\tau e^{-(x/\theta)^\tau}, \quad \text{for } x > 0.$$

This distribution satisfies the following:

$$\mathbb{E}[X^k] = \theta^k \Gamma(1 + k/\tau), \quad \text{for any } k > -\tau.$$

and

$$\mathbb{E}[(X \wedge x)^k] = \theta^k \Gamma(1 + k/\tau) \Gamma[1 + k/\tau; (x/\theta)^\tau] + x^k e^{-(x/\theta)^\tau}, \quad \text{for any } k > -\tau.$$

A discrete random variable  $N$  is said to belong to the  $(a, b, 0)$  class of distributions if it satisfies the relation

$$\Pr(N = k) = p_k = \left(a + \frac{b}{k}\right) \cdot p_{k-1}, \quad \text{for } k = 1, 2, \dots,$$

for some constants  $a$  and  $b$ . Alternatively, this relation can be expressed as a linear function given by

$$k \cdot \frac{p_k}{p_{k-1}} = b + ak, \quad \text{for } k = 1, 2, \dots$$

The initial value  $p_0$  is determined so that  $\sum_{k=0}^{\infty} p_k = 1$ .

The Poisson distribution  $\mathcal{PN}(\lambda)$  belongs to the  $(a, b, 0)$  class of distributions with  $a = 0$  and  $b = \lambda$ . Its mean and variance are:

$$\mathbb{E}[X] = \lambda \quad \text{and} \quad \text{Var}(X) = \lambda.$$